# WORDS CODING SET PARTITIONS

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ABSTRACT. The words in the title are characterized by the fact that a smaller number must (first) appear earlier than a larger number, and that all numbers  $1, \ldots, k$  are present (for some k). Under the assumption that the letters are drawn from a geometric distribution, the probability that a word of length n enjoys these properties is determined, both exactly and asymptotically.

### 1. INTRODUCTION

For a set partition of  $\{1, 2, ..., n\}$  into k blocks, a natural coding is as follows: Element 1 is in block 1, and the smallest number not in block 1 is in block 2, and the smallest number not in blocks 1 or 2, is in block 3, etc. In this way, to every element i a number  $a_i$  is attached, namely the block in which it lies. Writing these numbers as a word  $a_1 ... a_n$ , the set partition is coded in a natural way. One particular reference for this is [3].

Forgetting now about set partitions, we are talking about words where the letters are the positive integers, and, assuming that k is the largest letter that appears in the word, then the letters  $1, \ldots, k-1$  must also appear, and the word has exactly k (strict) left-to-right maxima, which is the same as saying that, if i < j, the first appearance of i is earlier than the first appearance of j. As one referee has kindly pointed out, such words are known as restricted growth strings in the literature [6].

Now we assign the (geometric) probability  $pq^{i-1}$  (where p + q = 1) to the letter *i* and consider  $P_n$ , the probability that a random word of length *n* has the *restricted growth* property. We are thus in the context of *combinatorics of geometrically distributed words*, a series of papers started with [4] and continued by the second writer as well as many others; a recent contribution is the paper [5].

The present question is not only appealing from a combinatorial point of view (easy to formulate but not trivial to solve) but the approach used here (with the parameter q) leads to "richer" results, and often the instance q = 1 corresponds to the classical combinatorial instance, especially, when the parameter is of the *order statistics* type.

We will prove the following theorems.

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**Theorem 1.** The probability  $P_n$  that a random word of length n has the restricted growth property is (exactly) given by

$$P_n = p \sum_{j=0}^{n-1} (-1)^j \binom{n-1}{j} q^j (p;q)_j.$$

Here we use the (standard) notation  $(x;q)_m = (1-x)(1-xq)\dots(1-xq^{m-1})$ . We will also need the limit of it as  $m \to \infty$ , denoted by  $(x;q)_{\infty}$ , as well as the Gaussian q-binomial coefficients

$$\begin{bmatrix} n \\ k \end{bmatrix}_q := \frac{(q;q)_n}{(q;q)_k (q;q)_{n-k}}$$

We need the following standard formulæ:

$$\sum_{k=0}^{N} {N \brack k}_{q} (-1)^{k} q^{\binom{k}{2}} x^{k} = (x;q)_{N},$$
$$\frac{1}{(w;q)_{\infty}} = \sum_{n \ge 0} \frac{w^{n}}{(q;q)_{n}}.$$

All this can be found in [1].

The asymptotic evaluation leads to our second theorem.

**Theorem 2.** The probability that a random word of length n has the restricted growth property is asymptotically given by

$$P_n \sim \frac{(p;q)_{\infty}}{L(q;q)_{\infty}} \Gamma\left(\frac{\log p}{\log q}\right) n^{-\frac{\log p}{\log q}} + n^{-\frac{\log p}{\log q}} \Phi(\log_Q n),$$

where  $\Phi(x)$  is a 1-periodic function with mean zero. The abbreviations Q = 1/q and  $L = \log Q$  are used. The function is given by its Fourier series

$$\Phi(x) = \frac{(p;q)_{\infty}}{L(q;q)_{\infty}} \sum_{k \neq 0} \Gamma\left(\frac{\log p}{\log q} + \frac{2\pi i k}{L}\right) e^{-2\pi i k x}.$$

In the symmetric case p = q, this looks better:

$$\frac{1}{L}n^{-1} + n^{-1}\Phi(\log_2 n).$$

## 2. Analysis

We use the natural decomposition

$$1\{\leq 1\}^* 2\{\leq 2\}^* 3\{\leq 3\}^* \dots k\{\leq k\}^*,$$

which translates into

$$\frac{zp}{1-(1-q)z}\frac{zpq}{1-(1-q^2)z}\dots\frac{zpq^{k-1}}{1-(1-q^k)z} = z^k p^k q^{\binom{k}{2}} \prod_{j=1}^k \frac{1}{1-(1-q^j)z}.$$

This has to be summed over all k, to get the generating function of the sought probabilities  $(P_n \text{ is the coefficient of } z^n \text{ in this series})$ :

$$\sum_{k \ge 1} z^k p^k q^{\binom{k}{2}} \prod_{j=1}^k \frac{1}{1 - (1 - q^j)z}.$$

Substituting z = w/(w-1), this becomes

$$\sum_{k\geq 1} w^k (-1)^k p^k q^{\binom{k}{2}} \prod_{j=1}^k \frac{1}{1-wq^j} = \sum_{k\geq 1} \frac{w^k (-1)^k p^k q^{\binom{k}{2}}}{(wq;q)_k}.$$

Reading off coefficients:

$$\begin{split} P_n &= [z^n] \sum_{k \ge 1} \frac{w^k (-1)^k p^k q^{\binom{k}{2}}}{(wq;q)_k} \\ &= \frac{1}{2\pi i} \oint \sum_{k \ge 1} \frac{dz}{z^{n+1}} \frac{w^k (-1)^k p^k q^{\binom{k}{2}}}{(wq;q)_k} \quad \text{by Cauchy's integral formula} \\ &= \frac{1}{2\pi i} \oint \sum_{k \ge 1} \frac{dw (1-w)^{n-1}}{w^{n+1}} \frac{w^k (-1)^{n-k} p^k q^{\binom{k}{2}}}{(wq;q)_k} \\ &= \sum_{k=1}^n [w^{n-k}] (1-w)^{n-1} \frac{(-1)^{n-k} p^k q^{\binom{k}{2}}}{(wq;q)_k} \\ &= \sum_{k=1}^n \sum_{j=0}^{n-1} \binom{n-1}{j} (-1)^j [w^{n-k-j}] \frac{(-1)^{n-k} p^k q^{\binom{k}{2}}}{(wq;q)_k} \\ &= \sum_{k=1}^n \sum_{j=0}^{n-1} \binom{n-1}{j} (-1)^{n-k-j} p^k q^{\binom{k}{2}} \left[ \binom{n-j-1}{k-1} \right]_q q^{n-k-j} \quad \text{the known expansion} \\ &= p \sum_{j=0}^{n-1} \binom{n-1}{j} q^{n-j-1} (-1)^{n-j-1} \sum_{k=0}^{n-j-1} (-1)^k p^k q^{\binom{k}{2}} \left[ \binom{n-j-1}{k} \right]_q \\ &= p \sum_{j=0}^{n-1} \binom{n-1}{j} q^{n-j-1} (-1)^{n-j-1} (p;q)_{n-j-1} \quad \text{the sum is known} \\ &= p \sum_{j=0}^{n-1} \binom{n-1}{j} q^j (-1)^j (p;q)_j. \end{split}$$

Is there a more direct way to prove this formula?

Here is an example for n = 3; the words enjoying the restricted growth property are 111, 112, 121, 122, 123, and they appear with probabilities  $p^3$ ,  $p^3q$ ,  $p^3q$ ,  $p^3q^2$ ,  $p^3q^3$ . And

$$p^{3} + p^{3}q + p^{3}q + p^{3}q^{2} + p^{3}q^{3} = p \sum_{j=0}^{2} {\binom{2}{j}} q^{j} (-1)^{j} (p;q)_{j} = p \left(1 - 2q(1-p) + q^{2}(1-p)(1-pq)\right).$$

For the asymptotic evaluation, we use the following integral representation as in [2]:

$$p\sum_{j=0}^{n-1} (-1)^j \binom{n-1}{j} q^j(p;q)_j = \frac{-p}{2\pi i} \int_{\mathcal{C}} q^z(p;q)_z \frac{\Gamma(n)\Gamma(-z)}{\Gamma(n-z)} dz.$$

Here, C enclosed the poles  $0, 1, \ldots, n-1$  and no others, and the interpretation of  $(p; q)_z$  is

$$(p;q)_z = \frac{(p;q)_\infty}{(pq^z;q)_\infty}.$$

For the readers' convenience we note that  $n! = \Gamma(n+1)$ , and thus

$$\frac{\Gamma(n)\Gamma(-z)}{\Gamma(n-z)} = \frac{\Gamma(n)}{(n-z-1)(n-z-2)\cdots(-z)} = \frac{(-1)^n(n-1)!}{z(z-1)\cdots(z+1-n)}$$

Furthermore, the residue of this expression at z = k is

$$\frac{(-1)^n(n-1)!}{k(k-1)\cdots 1\cdot (-1)\cdots (k+1-n)} = \frac{(-1)^{k-1}(n-1)!}{k!(n-1-k)!}.$$

To get asymptotics, we extend the contour of integration and have to consider the residues at the extra poles of

$$\frac{pq^{z}(p;q)_{\infty}}{(1-pq^{z})(pq^{z+1};q)_{\infty}}\frac{\Gamma(n)\Gamma(-z)}{\Gamma(n-z)}.$$

The poles with largest real part leading to the dominant contribution are at

$$z = -\frac{\log p}{\log q} + \frac{2\pi i k}{\log q}, \quad \text{for} \quad k \in \mathbb{Z}.$$

For k = 0 we get the interesting term, and the others define a small fluctuation around this value. We find:

$$\frac{pq^{-\frac{\log p}{\log q} + \frac{2\pi ik}{\log q}}(p;q)_{\infty}}{L(pq^{1-\frac{\log p}{\log q} - \frac{2\pi ik}{L}};q)_{\infty}} \frac{\Gamma(n)\Gamma(\frac{\log p}{\log q} + \frac{2\pi ik}{L})}{\Gamma(n + \frac{\log p}{\log q} + \frac{2\pi ik}{L})} = \frac{(p;q)_{\infty}}{L(q^{1-\frac{2\pi ik}{L}};q)_{\infty}} \frac{\Gamma(n)\Gamma(\frac{\log p}{\log q} + \frac{2\pi ik}{L})}{\Gamma(n + \frac{\log p}{\log q} + \frac{2\pi ik}{L})} \\ \sim \frac{(p;q)_{\infty}\Gamma(\frac{\log p}{\log q} + \frac{2\pi ik}{L})}{L(q;q)_{\infty}} n^{-\frac{\log p}{\log q} - \frac{2\pi ik}{L}}.$$

The term k = 0 leads to

$$\frac{(p;q)_{\infty}\Gamma(\frac{\log p}{\log q})}{L(q;q)_{\infty}}n^{-\frac{\log p}{\log q}}$$
$$n^{-\frac{\log p}{\log q}}\Phi(\log_Q n),$$

and the other ones to

where  $\Phi(x)$  is a 1-periodic function with mean zero. Note that

$$pq^{-\frac{\log p}{\log q} + \frac{2\pi ik}{\log q}} = 1,$$

which was used in these computations.

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